

# AN EXPERIMENTAL COMPARISON OF COKRIGING OF REGIONALIZED COMPOSITIONAL DATA USING FOUR DIFFERENT METHODS. CASE STUDY: BAUXITES IN HUNGARY

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## ABSTRACT

An important problem in the geosciences is the estimation or prediction of regionalized compositions. In fact, it is usual to deal with data such as percentages, concentrations, ppm,...., and use them to estimate values in other locations. Compositional data have been regarded as difficult to work with because of the so-called constant sum constraint. Following Aitchison (1986), any meaningful statement about a composition can be expressed in terms of logratios, but those transformations, and their backtransformations, are not always easy to deal with. The aim of this paper is to compare results obtained applying different methodologies developed in geostatistics, with samples of compositional data from a bauxite deposit in Halimba II (Hungary). Firstly, a classical geostatistics study is done using raw data; secondly applying two wellknown transformations in compositional data analysis: additive logratio (ALR) and centered logratio (CLR); thirdly, the Fast Fourier Transform (FFT) methodology to calculate the spatial variance-covariance matrix is used in cokriging. To be able to compare predictive values and kriging errors respective backtransforms are found. At last, results obtained with the different approaches are discussed and compared.

## 1. Introduction

The most common goal in Geostatistics is to estimate the value of an unknown variable in a location using the information given by some samples in its surroundings. A new problem comes when compositional variables are studied, those variables are characterised by their constant sum; that is, variables summing up to one (proportions), summing up to 100 (percentages) and so on. Their main features have been studied and described by many authors (Aitchison, Barceló, Egozcue, Pawlowsky and others) and have settled some specific methodologies to work with them. Those methodologies set up some transformation of data; the best-known ones are average logratio (ALR) and centered logratio (CLR). Recently, Yao and Journel have found some way to calculate covariances matrix with Fast Fourier Transform (FFT). So, it seems sensible to apply all those methods, as well as the classical one used by most geologists to the same data to assess their applicability and results. This is the goal of this paper: the use of all those four methods, finding out their difficulties and comparing their results with some well-known data. The database is a set of compositional data from a bauxite deposit named Halimba, which is the largest one in Europe continuously mined since 1950. Gy.Bárdossy, Budapest, furnished the data.

## 2. The data set

The studied deposit is in Hungary (Europe) and it is limited by East 117.6 - 114.0; North 13.0 - 8.8 geographic coordinates in a topographic map. The deposit covers an area of more than 8 km<sup>2</sup>; Halimba II is the only sector in the deposit that is still under prospection. The database consists of 55 samples representing 55 boreholes, after getting off 3 incomplete samples. In these boreholes the thickness of bauxite varies from 0.8 to 36.1 m. Variables used are the following: X = Easting; Y = Northing; V<sub>1</sub> = Concentration of Al<sub>2</sub>O<sub>3</sub>; V<sub>2</sub> = Concentration of

SiO<sub>2</sub>; V<sub>3</sub> = Concentration of Fe<sub>2</sub>O<sub>3</sub>; V<sub>4</sub> = Concentration of TiO<sub>2</sub>; V<sub>5</sub> = Concentration of H<sub>2</sub>O; V<sub>6</sub> = Concentration of CaO; V<sub>7</sub> = Concentration of MgO; concentrations are in percent. The values of V<sub>1</sub> to V<sub>7</sub> represent weighted averages in each borehole taken from intervals of 0.5 to 1.0 m length. Full database and histograms of the variables can be found at our website; table 1 shows the descriptive statistics of data set.

Table 1. Descriptive statistics of data set.

	Range	Minimum	Maximum	Average	Standard deviation	Symmetry	Kurtosis
<b>V1</b>	8.3	49.9	58.2	54.569	2.234	- 0.647	- 0.558
<b>V2</b>	7.4	0.7	8.1	3.889	2.007	0.334	- 0.876
<b>V3</b>	7.4	20.4	27.8	23.698	1.898	0.523	- 0.096
<b>V4</b>	2.1	1.6	3.7	2.778	0.332	-0.683	2.773
<b>V5</b>	2.3	11.3	13.6	12.371	0.499	0.477	0.128
<b>V6</b>	2.7	0.1	2.8	0.536	0.545	2.232	5.543
<b>V7</b>	1.8	0.1	1.9	0.267	0.327	3.133	11.688

### 3. Raw data geostatistical analysis

This is a traditional method to estimate any regionalised variable in geostatistics. It consists of building up variograms for each variable and cross-variograms when there are more than one of them. Once experimental (cross)variograms have been built they must be modeled. The corresponding theoretical ones are used in (co)kriging system to estimate the values on a regular grid. Variograms for the seven variables have been calculated and modeled; a table with the full description of those models can be found in our website. Once all variograms were built, cokriging has been done using KB2D program from GSLIB (1998).

### 4. Geostatistical analysis considering variables as compositions: ALR transform

Classical applications of geostatistics are related to mapping the spatial distribution of the variables under study. They give emphasis to characterize the variogram model and use the kriging (error) variance as a measure of estimation accuracy. Nowadays, some problems have been reported with compositional data. Those problems have been studied by many authors (references [1], [2], [3] and [8]). The main problem when handling compositional data is the so-called constant sum ( $K$ ) constraint. Usually  $K = 1$  or  $K = 100$ , if data are percentages. So, if  $V_1, \dots, V_N$  are proportions of  $N$  elements, then  $V_1 + \dots + V_N = K$ , which means that variables are not independent. To deal with compositional data and avoid this constraint, Aitchison has proposed some transforms. We have used two of them: average logratio (ALR) and centered logratio (CLR). With those transformations variables become independent and then classical kriging can be performed. As it is said beforehand,  $V_1, \dots, V_N$  must follow the constant sum constraint, but this quite never is true. Actually, we must define a new variable (called the residual) as  $V_R = K - (V_1 + \dots + V_N)$ . Then, the ALR variables  $U_i$  ( $i = 1, 2, \dots, N$ ) are defined as follows:

$$U_i = alr(V_i) = \log \frac{V_i}{V_R}, i = 1, 2, \dots, N \quad (1)$$

So now we are working with several  $U_i$  variables which do not follow the sum constraint; so they can be used as any other geostatistical data. We build and model their variograms (they can be found in our website). Once variogram was built, kriging has been done using KB2D

program from GSLIB (1998). Kriging results must be backtransformed to have the estimation of  $V_i$  in the grid; in this case, the corresponding ALR-backtransform is:

$$V_i = \frac{\exp U_i}{1 + \sum_{i=1}^N \exp U_i} K, i = 1, 2, \dots, N \quad (2)$$

### 5. Geostatistical analysis considering variables as compositions: CLR transform

Once  $V_R$  has been defined, we define a new variable as the geometrical mean of all of them:

$$V_{gm} = (V_1 V_2 \dots V_N V_R)^{\frac{1}{N+1}} = \exp \left( \frac{1}{N+1} \left( \sum_{i=1}^N (\ln V_i + \ln V_R) \right) \right) \quad (3)$$

Then, CLR transform consists in stating  $N+1$  new variables as:

$$W_j = clr(V_j) = \ln \frac{V_j}{V_{gm}}, j = 1, 2, \dots, N+1 \quad (4)$$

These  $N+1$  variables are not constrained, so they can be modeled and estimated. Once variograms have been built (they can be found in our website) cokriging has been done using KB2D program from GSLIB (1998). Then, backtransforms must be done to recover original variables. CLR-backtransform is:

$$V_j = \frac{\exp W_j}{\sum_{i=1}^{N+1} \exp W_i}, j = 1, 2, \dots, N+1 \quad (5)$$

### 6. Fast Fourier Transform method to calculate the covariance matrix

On the other hand, to avoid the modeling of variograms and crossvariograms, which may be very subjective, Yao and Journel (1998) have developed the so-called FFT method, which can be applied, in principle, to any kind of data. With FFT you do not need the independence of the variables and it builds up the covariance matrix, which can be used directly to kriging. This approach works as follows:

- a) Generate an experimental correlogram map on a regular grid. The grid typically has multiple nodes without estimates. The user has to specify the minimum number of data to be considered in the estimates at every node. This task is performed by program CORRMAP (see reference [6]).
- b) Program INTMAP fills in the blanks typically present in the grid generated in step 1 by using a smooth local interpolation.
- c) Program MULTSMTH corrects the smoothed grid to generate a third grid that is a tabulation of a positively semidefined correlogram. This condition is required to assure a unique solution for the kriging system of equations yielding a non-negative kriging variance.
- d) Convert the correlogram tabulation in step 3 to covariance tabulation by multiplying the correlogram grid by the sampling variance.
- e) As it was not possible to use KB2D to kriging, because with this method we obtain the covariance matrix and not the variograms, we had to change it (see reference [5]).

## 7. Results and discussion

Table 2 shows descriptive statistics for the estimations. Variables shown are raw estimations ( $V_i$ ), backtransformations of ALR estimations (BACK  $U_i$ ), backtransformations of CLR estimations (BACK  $W_i$ ) and estimations using FFT (FFT  $V_i$ ). Table 3 shows descriptive statistics of their differences, that is the differences between raw estimations and each of the other estimations. Kriging errors can be found in website. Figures comparing kriging results for the seven variables can be found in website; as an example you can see hereafter, in figure 1, results for variable  $V_1$ . In this figure, (a) is referred to raw data, (b) to the backtransformation of ALR-variable, (c) to the backtransformation of CLR-variable and (d) to the FFT transformation method. Looking at the contour maps, no significant differences among the first three methods arise. However (d)-picture, the one belonging to FFT method, shows higher resolution. It seems that it is because this method is less subjective.

## 8. Conclusions.

Using the results of this study some conclusions can be built:

- a) Kriging results in Halimba II using the four methods are quite similar.
- b) As regarding to the kriging errors, comparison is not so easy because it is not true that the backtransform of ALR and CLR transformations belong to the same space as the data (this is why Martin et al. defined stress).
- c) FFT method seems to be the best one, because it is less subjective, more precise and, furthermore, it is the easiest method to use. However, this method does not take into account if data are compositional or not.

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## Acknowledgements.

Authors want to thank specially:

- To Department of Applied Mathematics III, UPC.
- To the School of Civil Engineering (ETSECCPB) in UPC.
- To the “Consorti Escola Industrial de Barcelona” (CEIB).

Table 2. Descriptive statistics of kriging estimations.

Variable	N	Average	Median	Std. Dev.	Minimum	Maximum
V <sub>1</sub>	845	54.62	55.20	2.16	49.90	58.20
BACK U <sub>1</sub>	845	54.68	55.19	2.14	49.78	58.10
BACK W <sub>1</sub>	845	54.66	55.18	2.14	49.88	58.17
FFT V <sub>1</sub>	578	54.62	55.30	2.24	49.90	58.20
V <sub>2</sub>	845	3.929	3.800	2.000	0.700	8.100
BACK U <sub>2</sub>	845	3.898	3.626	2.007	0.700	8.150
BACK W <sub>2</sub>	845	3.899	3.623	2.006	0.700	8.150
FFT V <sub>2</sub>	578	3.924	3.602	2.011	0.700	8.100
V <sub>3</sub>	845	23.70	23.45	1.72	20.40	27.80
BACK U <sub>3</sub>	845	23.72	23.41	1.72	20.32	27.75
BACK W <sub>3</sub>	845	23.73	23.51	1.72	20.46	27.89
FFT V <sub>3</sub>	578	23.61	23.30	1.83	20.40	27.80
V <sub>4</sub>	845	2.777	2.800	0.348	1.600	3.700
BACK U <sub>4</sub>	845	2.779	2.814	0.348	1.600	3.710
BACK W <sub>4</sub>	845	2.779	2.813	0.346	1.600	3.680
FFT V <sub>4</sub>	578	2.783	2.801	0.341	1.600	3.700
V <sub>5</sub>	845	12.34	12.30	0.45	11.30	13.60
BACK U <sub>5</sub>	845	12.35	12.34	0.47	11.01	13.76
BACK W <sub>5</sub>	845	12.35	12.26	0.45	11.38	13.62
FFT V <sub>5</sub>	578	12.37	12.30	0.50	11.30	13.60
V <sub>6</sub>	845	0.504	0.300	0.465	0.100	2.800
BACK U <sub>6</sub>	845	0.490	0.301	0.453	0.090	2.810
BACK W <sub>6</sub>	845	0.490	0.301	0.453	0.100	2.810
FFT V <sub>6</sub>	578	0.503	0.300	0.489	0.100	2.800
V <sub>7</sub>	845	0.247	0.100	0.264	0.100	1.900
BACK U <sub>7</sub>	845	0.234	0.105	0.247	0.090	1.900
BACK W <sub>7</sub>	845	0.234	0.101	0.247	0.100	1.900
FFT V <sub>7</sub>	578	0.249	0.100	0.293	0.100	1.900

Table 3. Descriptive statistics of estimation differences.

	N	Average	Median	Std. Dev.	Minimum	Maximum
V <sub>1</sub> - BACK U <sub>1</sub>	845	-0.0621	-0.0088	0.2640	-2.6162	0.6762
V <sub>1</sub> - BACK W <sub>1</sub>	845	-0.0429	0.0009	0.2104	-1.7219	0.3092
V <sub>1</sub> - FFT V <sub>1</sub>	560	0.0189	0.0000	0.6570	-3.1083	6.7002
V <sub>2</sub> - BACK U <sub>2</sub>	845	0.0310	-0.0001	0.2165	-1.2449	2.0673
V <sub>2</sub> - BACK W <sub>2</sub>	845	0.0301	-0.0011	0.1993	-1.1678	1.7569
V <sub>2</sub> - FFT V <sub>2</sub>	560	0.0026	0.0000	0.4389	-3.9004	2.8890
V <sub>3</sub> - BACK U <sub>3</sub>	845	-0.0148	-0.0121	0.1170	-0.9911	0.4930
V <sub>3</sub> - BACK W <sub>3</sub>	845	-0.0299	-0.0262	0.1234	-1.1340	0.7786
V <sub>3</sub> - FFT V <sub>3</sub>	560	0.0174	0.0000	0.6029	-3.5003	6.5000
V <sub>4</sub> - BACK U <sub>4</sub>	845	-0.0018	0.0023	0.0268	-0.2413	0.2638
V <sub>4</sub> - BACK W <sub>4</sub>	845	-0.0017	0.0005	0.0260	-0.1852	0.2491
V <sub>4</sub> - FFT V <sub>4</sub>	560	0.0047	0.0000	0.0667	-0.6930	0.4000
V <sub>5</sub> - BACK U <sub>5</sub>	845	-0.0044	0.0034	0.1598	-1.0208	1.2711
V <sub>5</sub> - BACK W <sub>6</sub>	845	-0.0072	-0.0039	0.0685	-0.6018	0.2735
V <sub>5</sub> - FFT V <sub>6</sub>	560	-0.0116	0.0000	0.1633	-1.0505	1.0498
V <sub>6</sub> - BACK U <sub>6</sub>	845	0.0138	-0.0001	0.0633	-0.0745	0.6584
V <sub>6</sub> - BACK W <sub>6</sub>	845	0.0142	0.0000	0.0681	-0.0285	0.7951
V <sub>6</sub> - FFT V <sub>6</sub>	560	-0.0054	0.0000	0.1530	-1.2472	1.2500
V <sub>7</sub> - BACK U <sub>7</sub>	845	0.0133	0.0001	0.0673	-0.0251	0.6202
V <sub>7</sub> - BACK W <sub>7</sub>	845	0.0135	0.0001	0.0666	-0.0057	0.5916
V <sub>7</sub> - FFT V <sub>7</sub>	560	-0.0040	0.0000	0.1144	-1.0331	0.9000

Figure 1. Contour maps of variable  $V_1$ .

